

Fighting Noise with Noise

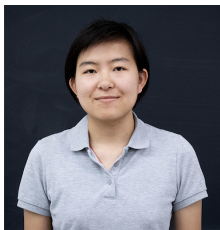
Causal Inference with Many Candidate Instruments

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University of Toronto

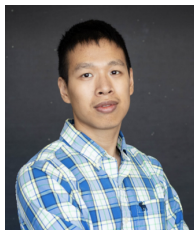


Pacific causal inference conference
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Acknowledgements



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Stanislav Volgushev



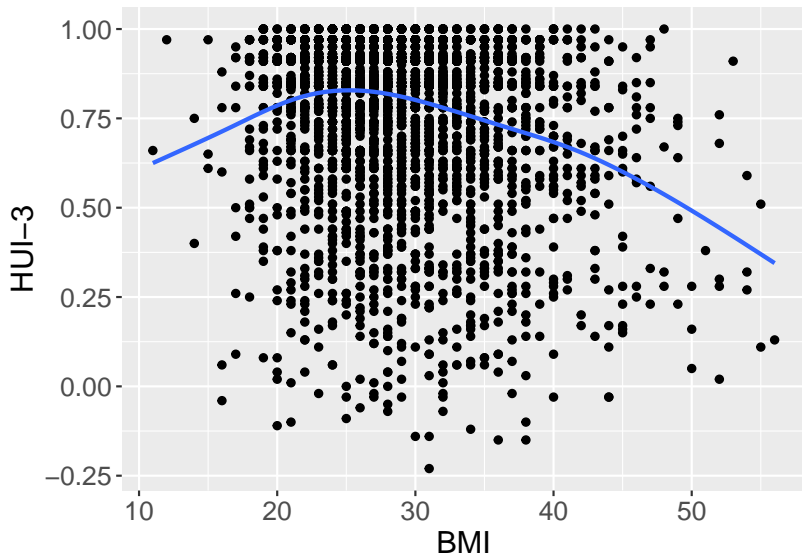
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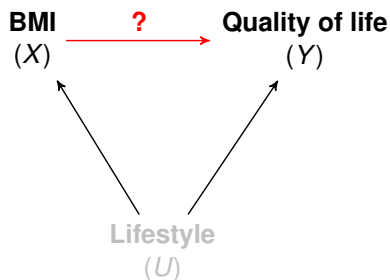
Overweight



Overweight is associated with Quality of Life

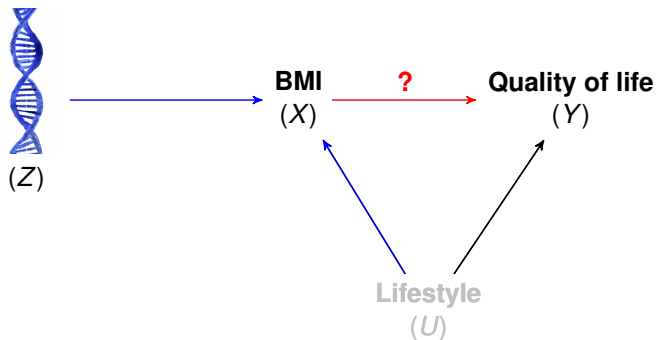


Is this association causal?



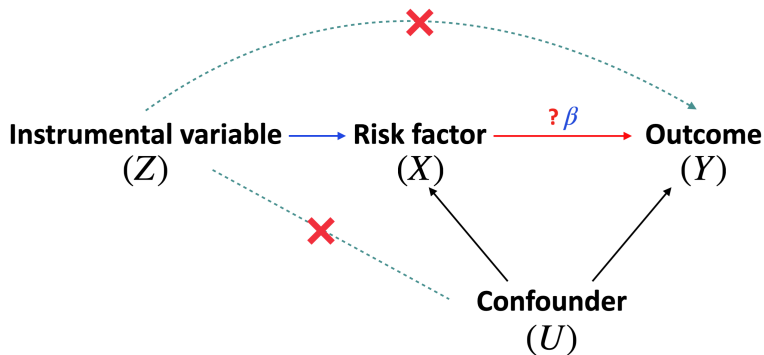
☛ **Unmeasured confounding!**

Mendelian randomization (MR)



Nature rolls the dice!

The math behind



$$\left. \begin{array}{l} \text{IV1. } Z \rightarrow X \neq 0 \\ \text{IV2. } Z \rightarrow Y = 0 \\ \text{IV3. } \text{cov}(Z, U) = 0 \end{array} \right\} \Rightarrow \begin{array}{l} Z \rightarrow X \rightarrow Y = (Z \rightarrow X) \times \underbrace{(X \rightarrow Y)}_{\beta} \\ X \rightarrow Y = \frac{Z \rightarrow X \rightarrow Y}{Z \rightarrow X} \end{array}$$

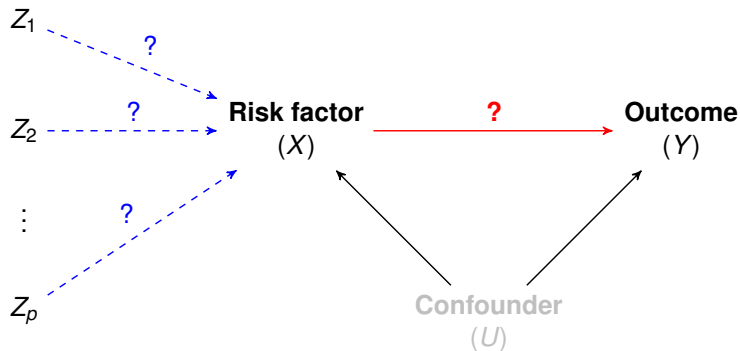
Natural experiments are imperfect!

1. Most genetic variants are **irrelevant!**
 - In the past: rely on expert knowledge

Natural experiments are imperfect!

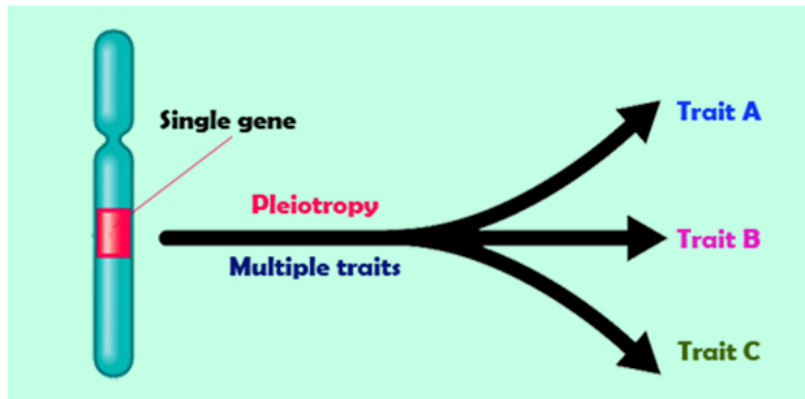
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Modern approach: GWAS



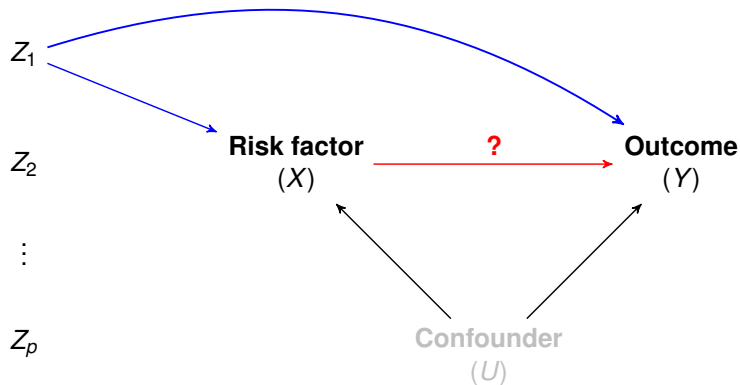
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- 2 Some of SNPs that pass the GWAS test may still be **invalid** due to pleiotropy



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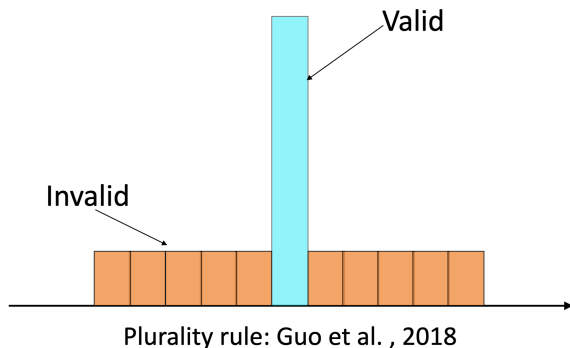
- 2 Some of SNPs that pass the GWAS test may still be **invalid** due to pleiotropy

Existing solution: Valid causal inference with some invalid instruments (Bowden et al. [2015], Kolesár et al. [2015], Kang et al. [2016], Hartwig et al. [2017], Guo et al. [2018], Hartford et al. [2020])

Valid causal inference with some invalid instruments

Solution:

*All happy instruments are alike;
each unhappy instrument is unhappy in its own way*



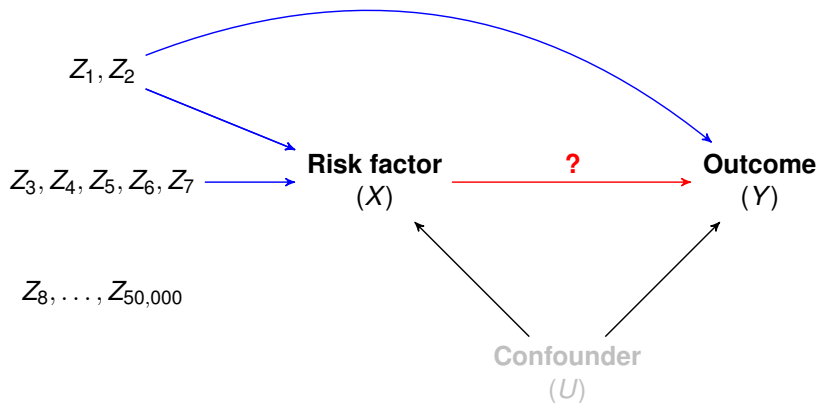
Summary of existing solutions

Step 1 GWAS based on strength of the links $Z_j \rightarrow X, j = 1, \dots, p$

Step 2 Apply a mode-finding algorithm to identify the valid IVs

Step 3 Use the valid IVs to estimate the causal effect

A toy simulation



- Sample size = 500
- $Z_j, j = 1, \dots, 50,000, U \sim N(0, 1)$
- Linear models
- True causal effect $X \rightarrow Y$ is 2

Step 1: GWAS

Step 1.1 (Marginal screening): Select the top 500 candidate IVs Z_r based on the marginal correlation

$$|Cor(Z_j, X)|, j = 1, \dots, 50,000$$

Step 1.2 (Joint thresholding): Fit a debiased Lasso model

$$X \sim Z_r$$

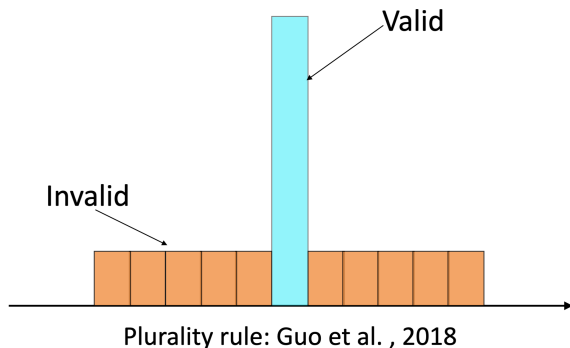
Result: On average left with 21 candidate IVs

😊 Include the 7 relevant IVs (5 valid, 2 invalid) every single time!

Step 2: Mode-finding

Solution:

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Step 3: Causal effect estimation

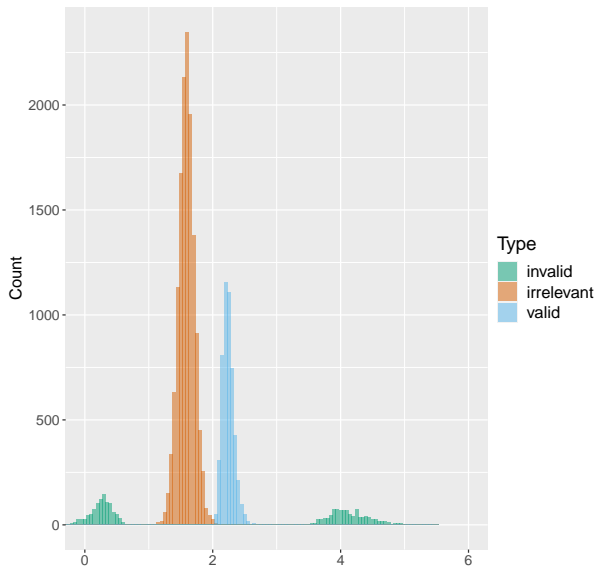
- ☹ The causal effect estimates are biased!
 - Bias = -0.22 (SE = 0.02)
 - Coverage probability = 0.07 (nominal = 0.95)

What went wrong?

Recall in Step 1: Select 21 candidate IVs

- 5 valid
- 2 invalid
- **14 irrelevant with spurious correlation**

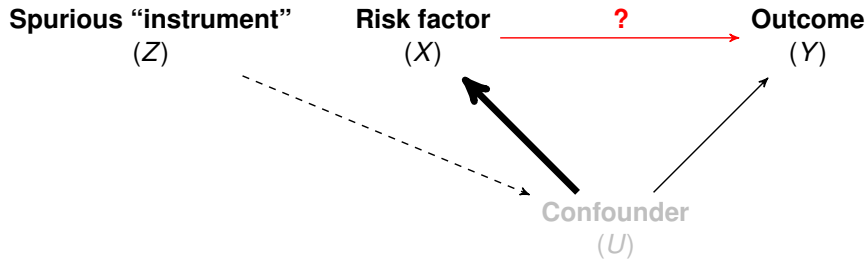
What happened in Step 2



Distribution of causal effect estimates for candidate IVs passing joint thresholding across 1000 Monte Carlo runs

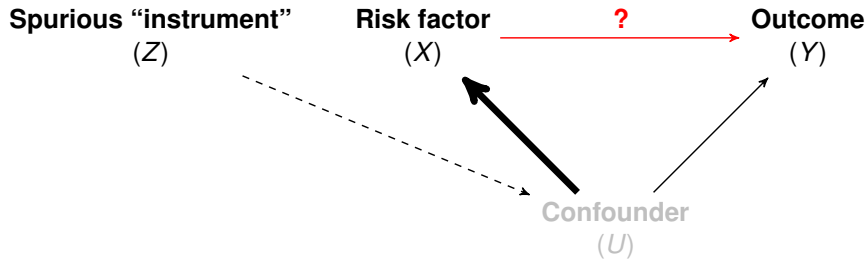
All irrelevant “instruments” are alike. Why?

The simplest case (extreme confounding)



- Single unmeasured confounder U
- X is only determined by U

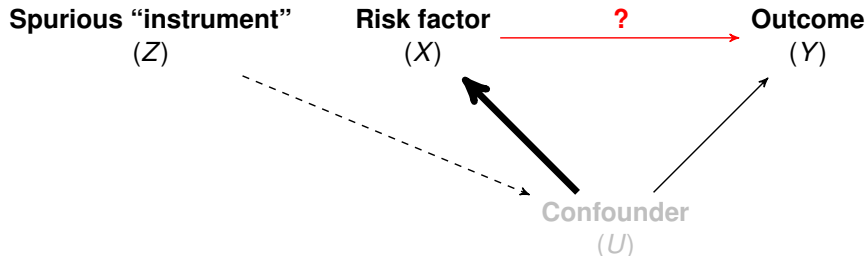
The simplest case (extreme confounding)



Recall that under the IV framework,

$$X \rightarrow Y = \frac{Z \rightarrow X \rightarrow Y}{Z \rightarrow X} = \frac{\text{Cov}(Z, Y)}{\text{Cov}(Z, X)}$$

The simplest case (extreme confounding)



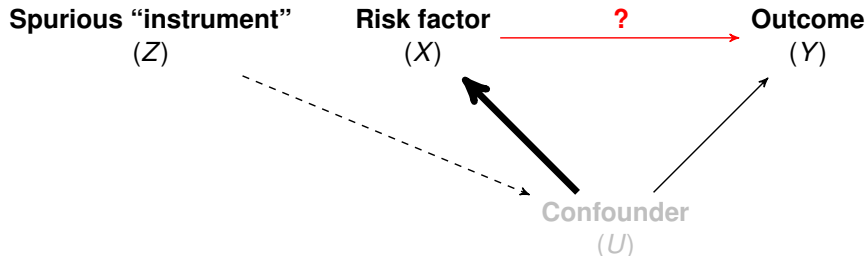
In the case of irrelevant "instrument,"

- $Cov(Z, Y) = Z \rightarrow U \rightarrow X \rightarrow Y + Z \rightarrow U \rightarrow Y$
- $Cov(Z, X) = Z \rightarrow U \rightarrow X$

$$\text{So } \frac{Cov(Z, Y)}{Cov(Z, X)} = X \rightarrow Y + \frac{U \rightarrow Y}{U \rightarrow X}$$

The bias is the same for all irrelevant "instruments"!

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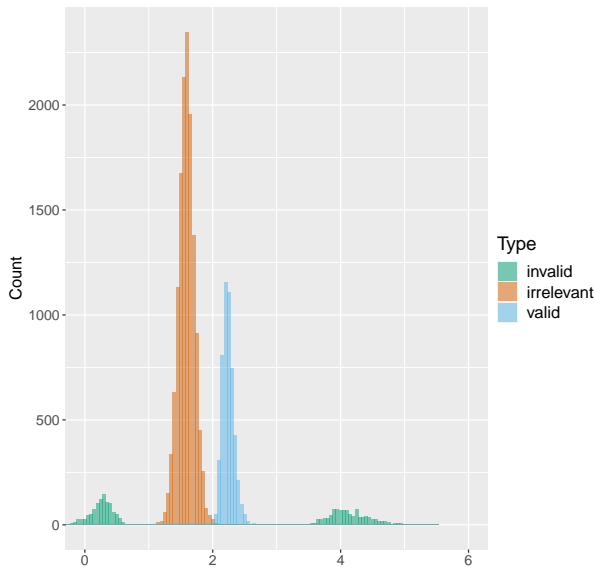
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$$\text{So } \frac{Cov(Z, Y)}{Cov(Z, X)} = X \rightarrow Y + \frac{U \rightarrow Y}{U \rightarrow X}$$

The bias is the same for all irrelevant "instruments"!

- In fact, also same as the bias from $Y \sim X$

What happened in Step 2



Distribution of causal effect estimates for candidate IVs passing joint thresholding across 1000 Monte Carlo runs

What we know so far:

- All **valid** instruments are alike
- All **irrelevant** instruments are alike
- **Invalid** instruments are different from each other

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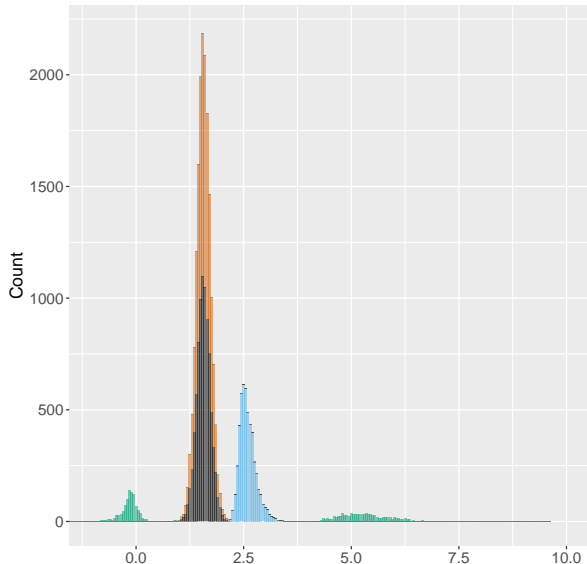
- All **valid** instruments are alike
- All **irrelevant** instruments are alike
- **Invalid** instruments are different from each other

How to distinguish valid instruments from irrelevant ones?

Key idea: Generate **pseudo variables** by permuting rows of Z , denoted by Z^*

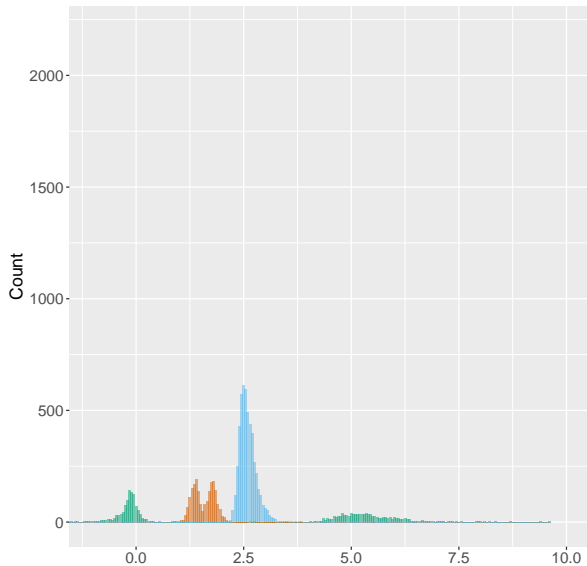
- **Group 1: valid instruments in Z**
- **Group 2: invalid instruments in Z**
- **Group 3: irrelevant instruments in Z**
- **Group 4: irrelevant instruments in Z^***

Groups 3 and 4 are alike! We can track variables in Group 4.



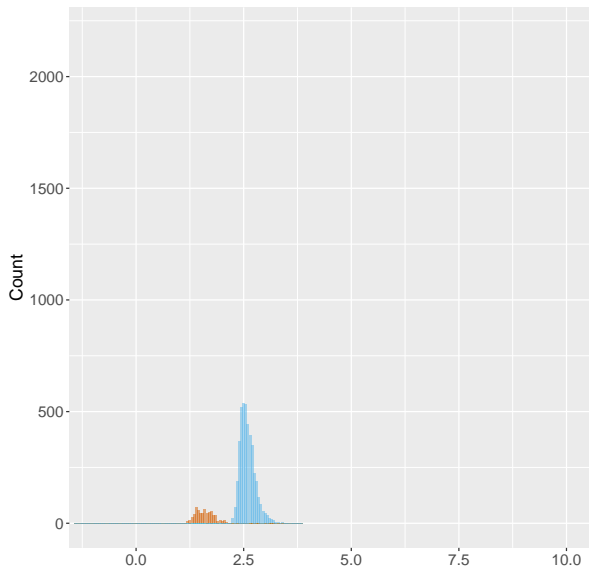
Distribution of causal effect estimates for candidate IVs passing joint thresholding across 1000 Monte Carlo runs

Step 2: Remove causal effect estimates inside the range of pseudo effect estimates



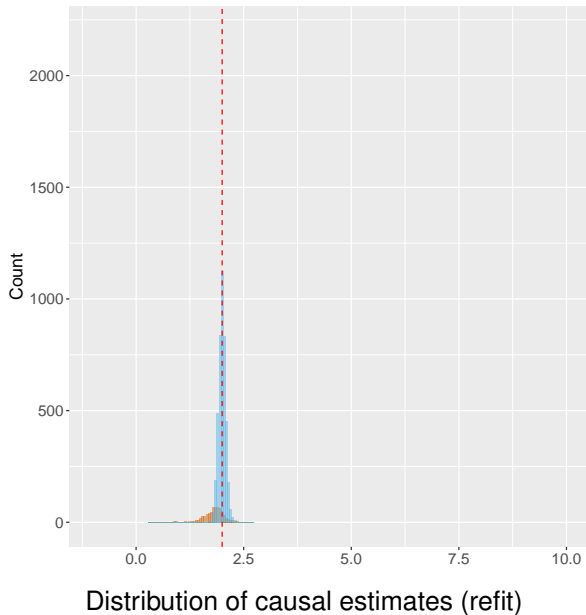
Distribution of causal estimates after removing spurious IVs

Step 3: Develop a mode-finding algorithms to identify the **valid** instruments



Distribution of causal estimates after mode-finding

Step 4: Use the selected valid instruments to estimate the causal effects



Theoretical guarantees

Theorem 1 (Zhang et al., 2022)

(Spurious IV) Under some regularity conditions, with probability tending to 1, causal estimates for spurious IVs are concentrated within

$$[\beta^* + C_* - d, \beta^* + C_* + d]$$

for $d > L(\omega, \sigma_X^2, \sigma_u^2, \sigma_Y^2, \alpha_X^*, \alpha_Y^*)$, where $C_* = (\alpha_Y^* \alpha_X^* \sigma_u^2) / (\alpha_X^{*2} \sigma_u^2 + \sigma_X^2)$.

Theorem 2 (Zhang et al., 2022)

(Valid IV) Under some regularity conditions, with probability tending to 1, causal estimates for valid IVs are concentrated around β^* .

Separation: If $d < |C_*|$, then spurious IVs and valid IVs are separable.

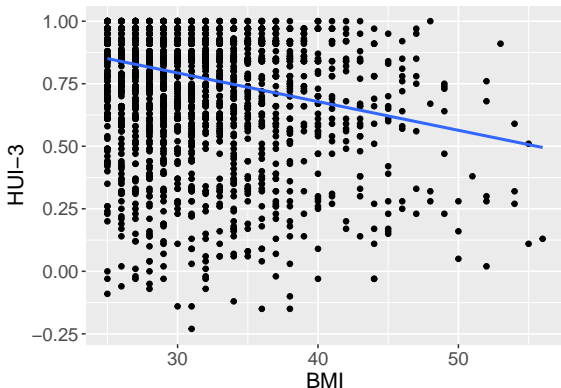
Easier separation if stronger U OR sufficiently large ω

Wisconsin Longitudinal Study



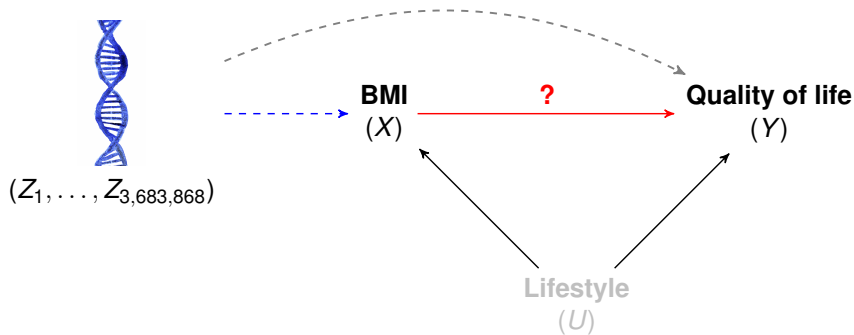
- Participants were graduates from Wisconsin high schools in 1957 and their siblings
- Use 3023 unrelated individuals re-interviewed in 2011
- Exposure: BMI (only those ≥ 25)
 - mean (SD): 30.6 (4.9)
- Outcome: Health Utility Index Mark 3 (HUI-3)
 - between -0.22 and 1
 - mean (SD): 0.79 (0.23)

Crude analysis



After adjusting for age, gender and education, one unit BMI increase associated with 0.011 unit decrease in HUI-3 (95% CI = [0.009, 0.013])

Conceptual Mendelian randomization model

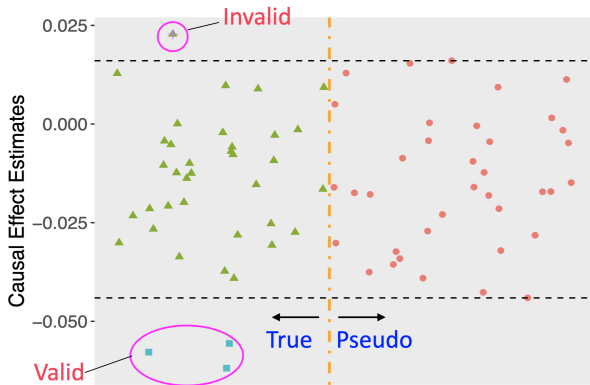


- Consider baseline covariates including age, gender, education and population stratification (top 6 principal components)

Data analysis: Preprocessing + Step 1

- Quality control → 3,683,868 SNPs left
- Generate the same number of pseudo SNPs by permuting the rows of the Z matrix
- After GWAS: 44 SNPs + 42 pseudo

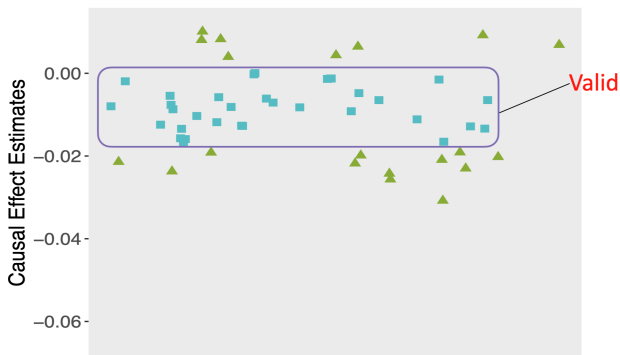
Data analysis: Steps 2-4



Causal estimate: one unit increase in BMI will result in **0.039** unit decrease of HUI-3 (95% CI = $[-0.052, -0.025]$)

Data analysis: Ignoring spurious instruments

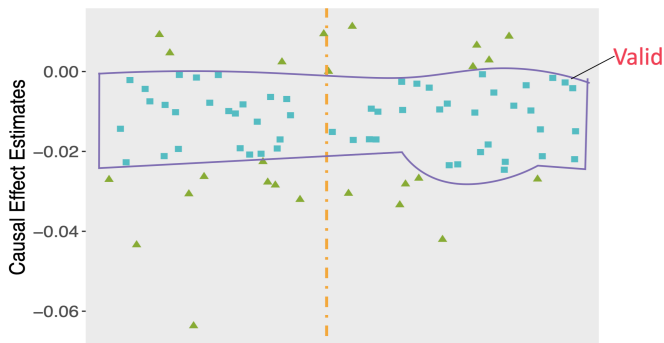
Two-stage hard thresholding with voting (Guo et al., 2018): true SNPs only



Causal estimate: one unit increase in BMI will lead to **0.008** unit decrease in HUI-3 (95% CI = [-0.013, -0.003])

Data analysis: Ignoring spurious instruments

Two-stage hard thresholding with voting (Guo et al., 2018): same data set as the proposed (true + pseudo)



Causal estimate: one unit increase in BMI will lead to **0.011** unit decrease in HUI-3 (95% CI = $[-0.015, -0.008]$)

Comparison of results

Proposed:

1 SD increase in BMI leads to **1 SD** (roughly) decrease in HUI-3

Guo et al. (2018)'s method (ignore spurious instruments; true only):

1 SD increase in BMI leads to **0.2 SD** (roughly) decrease in HUI-3

Guo et al. (2018)'s method (ignore spurious instruments; true + pseudo):

1 SD increase in BMI leads to **0.2 SD** (roughly) decrease in HUI-3

Crude analysis (OLS):

1 SD increase in BMI is associated with **0.2 SD** (roughly) decrease in HUI-3

Summary

- Mendelian randomization is a powerful tool for causal effect estimation
- Challenges for MR studies
 - Find relevant IVs → GWAS
 - Deal with pleiotropy → find the mode of effect estimates
 - **Bias due to spurious variables in GWAS**
- Spurious IV bias is potentially a serious problem
 - Ignoring Spurious IV bias has a similar effect as ignoring unmeasured confounding bias
- **Fight noise with noise**
 - Use pseudo variables to correct for Spurious IV bias

References I

- Guo, Z., Kang, H., Cai, T. T. & Small D. S. (2018). Confidence intervals for causal effects with invalid instruments by using two-stage hard thresholding with voting. *J. R. Statist. Soc. B.* 80, 793–815.
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